
On Linear Innovations Sequence

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1 PROBLEM STATEMENT

Suppose X, Y_1, Y_2, \dots, Y_n are random vectors and $1, Y_1, Y_2, \dots, Y_n$ have finite second moments but not orthogonal. According to [1], by orthogonalizing this sequences we can obtain a new sequence $1, \tilde{Y}_1, \tilde{Y}_2, \dots, \tilde{Y}_n$ used for linear estimating of X . Let $\tilde{Y}_1 = Y_1 - E[Y_1]$, and for $i \geq 2$, let

$$\tilde{Y}_i = Y_i - \hat{E}[Y_i | Y_1, Y_2, \dots, Y_{i-1}] \quad (1.1)$$

It is claimed in the following elaboration (online version of [1], page 95 and 96) that

1. $E[\tilde{Y}_i] = 0$ for all i and $E[\tilde{Y}_i \tilde{Y}_j^T] = 0$ for all $i \neq j$.
2. $\hat{E}[X | Y_1, Y_2, \dots, Y_n] = \hat{E}[X | \tilde{Y}_1, \tilde{Y}_2, \dots, \tilde{Y}_n]$.

2 ELABORATION

First, prove $E[\tilde{Y}_i] = 0$ for all i . $E[\tilde{Y}_1] = E[Y_1] - E[E[Y_1]] = 0$. Denote $Z_i^T = [Y_1^T, Y_2^T, \dots, Y_{i-1}^T]$. For $i \geq 2$,

$$\hat{E}[Y_i | Y_1, Y_2, \dots, Y_{i-1}] = E[Y_i] + \text{Cov}[Y_i, Z_i] \text{Cov}[Z_i]^{-1} (Z_i - E[Z_i]). \quad (2.1)$$

Apparently, $E[\hat{E}[Y_i | Y_1, Y_2, \dots, Y_{i-1}]] = E[Y_i]$. Therefore, $E[\tilde{Y}_i] = 0$ for $i \geq 2$ also holds. Then prove $E[\tilde{Y}_i \tilde{Y}_j^T] = 0$ for all $i \neq j$. Without loss of generality, suppose that $i > j$. According to [1] page 86, $\tilde{Y}_i = e_i$ and is orthogonal to all linear combination of $1, Y_1, Y_2, \dots, Y_{i-1}$. \tilde{Y}_j is a linear combination of $1, Y_1, Y_2, \dots, Y_j$. Since $i > j$, \tilde{Y}_j is also a linear combination of $1, Y_1, Y_2, \dots, Y_{i-1}$. Therefore $E[\tilde{Y}_i \tilde{Y}_j^T] = 0$ for all $i \neq j$.

Second, according to Eq. (1.1), \tilde{Y}_i is a linear function of $1, Y_1, Y_2, \dots, Y_i$. Therefore we can claim that

$$\tilde{Y} = BY + c \quad (2.2)$$

where matrix B is lower triangular matrix with all diagonal elements to be 1. Therefore B is invertible. By [1],

$$\hat{E}[X|Y_1, Y_2, \dots, Y_n] = E[X] + \text{Cov}[X, Y]\text{Cov}[Y]^{-1}(Y - E[Y]) \quad (2.3)$$

and similarly

$$\hat{E}[X|\tilde{Y}_1, \tilde{Y}_2, \dots, \tilde{Y}_n] = E[X] + \text{Cov}[X, \tilde{Y}]\text{Cov}[\tilde{Y}]^{-1}(\tilde{Y} - E[\tilde{Y}]) \quad (2.4)$$

Since $\text{Cov}[\tilde{Y}] = B\text{Cov}[Y]B^T$ which leads to

$$\text{Cov}[\tilde{Y}]^{-1} = (B^T)^{-1}\text{Cov}[Y]^{-1}B^{-1}$$

and

$$\text{Cov}[X, \tilde{Y}] = \text{Cov}[X, BY + c] = \text{Cov}[X, Y]B^T,$$

Eq. (2.4) can be reformulated as

$$\hat{E}[X|\tilde{Y}_1, \tilde{Y}_2, \dots, \tilde{Y}_n] = E[X] + \text{Cov}[X, Y]B^T(B^T)^{-1}\text{Cov}[Y]^{-1}B^{-1}B(Y - E[Y]) = \hat{E}[X|Y_1, Y_2, \dots, Y_n] \quad (2.5)$$

REFERENCES

- [1] Bruce Hajek, *Random processes for engineers*, Cambridge University Press, 2015.